

Coarse categorical data under ontologic and epistemic uncertainty

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08th of September 2014

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- 2 Data under ontologic uncertainty
 - Motivation
 - Basic idea of analysis
- 3 Data under epistemic uncertainty
 - Motivation
 - Analysis of two different situations
- 4 Summary

Epistemic vs. ontic/ontologic uncertainty (I. Couso, D. Dubois, 2014)

► Ontologic uncertainty

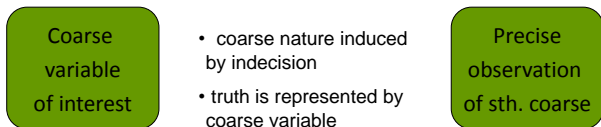
Coarse
variable
of interest

- coarse nature induced by indecision
- truth is represented by coarse variable

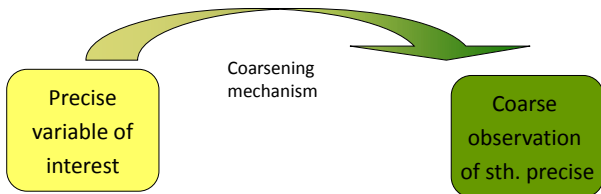
Precise
observation
of sth. coarse

Epistemic vs. ontic/ontologic uncertainty (I. Couso, D. Dubois, 2014)

▶ Ontologic uncertainty



▶ Epistemic uncertainty



Why should data under ontologic uncertainty be collected?

Example:

Which party will you give your vote?

A

B

C

Don't know

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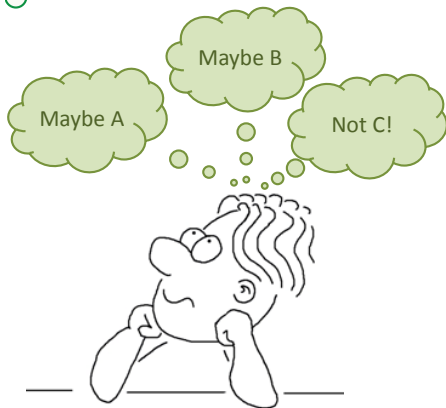
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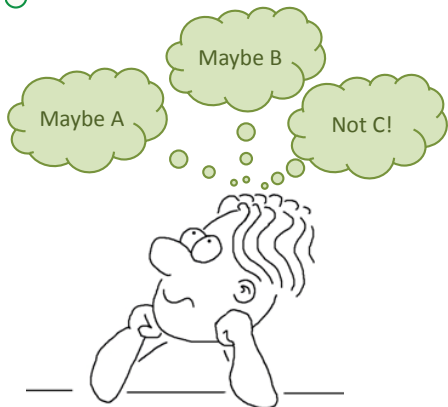
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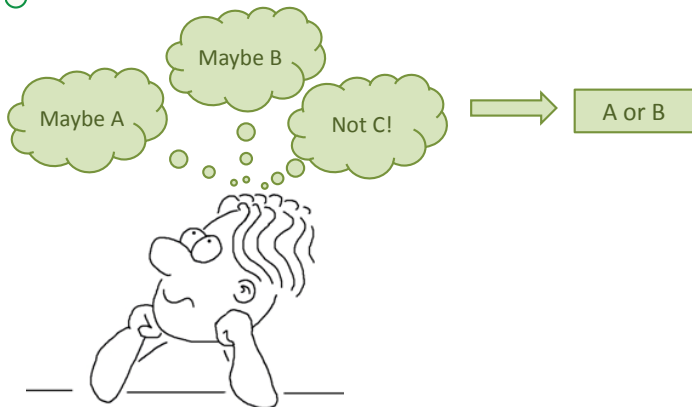
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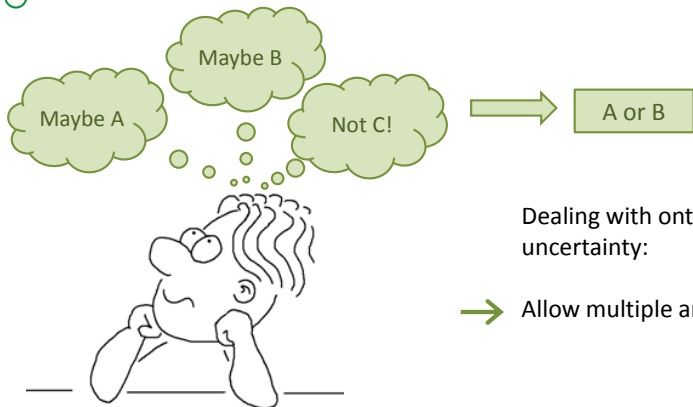
B



C



Don't know



Basic idea (illustrated by GLES 2013)

certainty	vote	assesCD	assesSPD	assesGREEN	assesLEFT	...
very certain	SPD	-1	5	2	1	...
certain	CD	4	3	3	1	...
not that certain	GREEN	3	4	4	-1	...
not certain at all	CD	-3	2	2	2	...

Exemplary
extraction of
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Analysis:

Prediction

- Classical analysis - Vote of respondents who are certain:

Prediction for "CD": $\frac{\# \text{respondents who will vote for "CD" with certainty}}{\# \text{respondents who are certain}} = 0.46$

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- Dealing with ontologic uncertainty:

CD 519	SPD 287	GREEN 105	LEFT 101	OTHER 62
SPD-CD 34	GREEN-SPD 34	CD-OTHER 24	LEFT-SPD 14	GREEN-LEFT 13
GREEN-SPD-CD 13	SPD-CD-OTHER 13	LEFT-GREEN-SPD 13	SPD-OTHER 12	rare comb. 77

$$\hat{B}_{el}(CD) = \frac{519}{1321} = 0.39$$

$$\hat{P}(CD) = \frac{519 + 34 + 24 + 13 + 13}{1321} = 0.45$$

⇒

Prediction for "CD":
[0.39, 0.45]



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Exemplary
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Analysis:

Regression

- For reasons of explanation interpretation of selected β estimators only
- Reference category: "SPD"
- Classical analysis - Vote of respondents who are certain:

	Intercept	sexFEM	InfoNEWSPAPER	InfoRADIO	InfoWEB
CD	0.2914	0.2456	-0.0037	-0.1487	-0.6774

Model under ontologic uncertainty

Data under ontologic uncertainty:

- Y_i : categorical random variable of nominal scale of measurement with $Y_i \subseteq \underbrace{\{a, b, \dots\}}_{\text{precise categories}}$
- $m = |\mathcal{P}(\Omega) \setminus \emptyset|$: number of categories of Y_i

Model under ontologic uncertainty:

⇒ classical multinomial logit model with different number of categories:

The probability of occurrence for category $r = 1, 2, 3, \dots, m - 1$ can be calculated by

$$P(Y_i = r | \mathbf{x}_i) = \frac{\exp(\mathbf{x}_i^T \beta_r)}{1 + \sum_{s=1}^{m-1} \exp(\mathbf{x}_i^T \beta_s)}$$

and for category m by

$$P(Y_i = m | \mathbf{x}_i) = \frac{1}{1 + \sum_{s=1}^{m-1} \exp(\mathbf{x}_i^T \beta_s)}$$

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- Dealing with ontologic uncertainty:

	Intercept	sexFEM	InfoNEWSPAPER	InfoRADIO	InfoWEB
CD	0.4706	0.3135	-0.0784	0.1856	-0.5025
SPD-CD	-2.2007	0.4034	-1.2556	0.9476	0.1886
GREEN-SPD	-2.4432	0.9393	-1.3053	0.2300	0.1844

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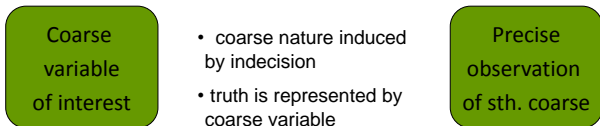
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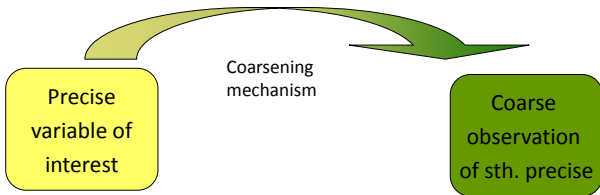
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Epistemic vs. ontologic uncertainty

▶ Ontologic uncertainty



▶ Epistemic uncertainty



When do data under epistemic uncertainty occur?

Reasons for coarse categorical data:

- Guarantee of anonymization, prevention of refusals

Example:

“Which kind of party did you elect?”

rather left center rather right

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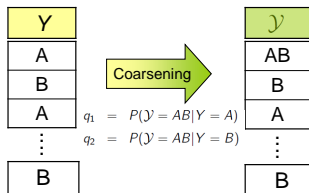
rather left center rather right

- Different levels of reporting accuracy
(lack of knowledge, vague question formulation)

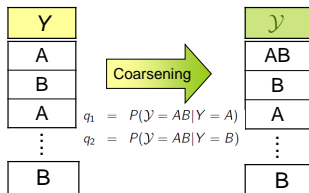
Examples:

“Which car do you drive?”

Addressed data situations



Addressed data situations



Two different situations will be regarded:

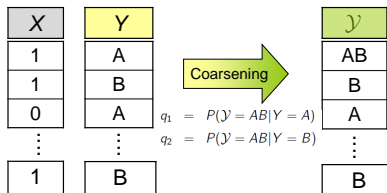
Case 1 - No covariates:

- IID-assumption

$$\pi_{iA} = \pi_A$$

- Constant coarsening mechanisms q_1 and q_2

Addressed data situations



Two different situations will be regarded:

Case 1 - No covariates:

- IID-assumption

$$\pi_{iA} = \pi_A$$

- Constant coarsening mechanisms q_1 and q_2

Case 2 - Binary covariates, no intercept

$$\pi_{iA|X_i=1} = \frac{\exp(\beta_A)}{1 + \exp(\beta_A)}$$

$$\pi_{iA|X_i=0} = \frac{1}{2}$$

- Constant coarsening mechanisms q_1 and q_2

Case 1 - No covariates + IID-assumption

log-Likelihood under the iid assumption :

$$\begin{aligned}
 l(\pi_A, q_1, q_2) &= \ln \left(\prod_{i:Y_i=A} \underbrace{P(Y=A|Y=A)}_{(1-q_1)} \pi_{iA} \prod_{i:Y_i=B} \underbrace{P(Y=B|Y=B)}_{(1-q_2)} (1-\pi_{iA}) \right. \\
 &\quad \left. \prod_{i:Y_i=AB} \underbrace{P(Y=AB|Y=A)}_{q_1} \pi_{iA} + \underbrace{P(Y=AB|Y=B)}_{q_2} (1-\pi_{iA}) \right) \\
 &\stackrel{iid}{=} n_A \cdot [\ln(1-q_1) + \ln(\pi_A)] + n_B \cdot [\ln(1-q_2) + \ln(1-\pi_A)] \\
 &\quad n_{AB} \cdot [q_1 \pi_A + q_2 (1-\pi_A)]
 \end{aligned}$$

FOC:

$$\begin{aligned}
 \text{I.) } \frac{\partial}{\partial \pi_A} &= \frac{n_{AB}}{q_1 \pi_A + q_2 (1-\pi_A)} (q_1 - q_2) + \frac{n_A}{\pi_A} - \frac{n_B}{1-\pi_A} \stackrel{!}{=} 0 \\
 \text{II.) } \frac{\partial}{\partial q_1} &= \frac{n_{AB}}{q_1 \pi_A + q_2 (1-\pi_A)} \pi_A - \frac{n_A}{1-q_1} \stackrel{!}{=} 0 \\
 \text{III.) } \frac{\partial}{\partial q_2} &= \frac{n_{AB}}{q_1 \pi_A + q_2 (1-\pi_A)} (1-\pi_A) - \frac{n_B}{1-q_2} \stackrel{!}{=} 0
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 \end{aligned}$$

Necessary and sufficient condition for estimators $(\hat{\pi}_A, \hat{q}_1, \hat{q}_2)$

$$\frac{n_{AB}}{n} = \hat{q}_1 \cdot \hat{\pi}_A + \hat{q}_2 \cdot (1 - \hat{\pi}_A)$$

Case 2 - Binary covariates, no intercept

Involved assumptions:

- *No intercept β_0*

$$\Rightarrow \pi_{iA|X_i=1} = \frac{\exp(\beta_A)}{1+\exp(\beta_A)} \text{ and } \pi_{iA|X_i=0} = \frac{1}{2}$$

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- *Constant coarsening mechanisms q_1 and q_2 :*

Example: "Do you regularly steal candy (\boxtimes) out of your mother's candy box?"

Asked: girls (g) and boys (b)

X	Y	\mathcal{Y}
g	\boxtimes	$\bar{\boxtimes}$ or $\bar{\boxtimes}$
g	$\bar{\boxtimes}$	$\bar{\boxtimes}$
g	\boxtimes	\boxtimes
b	\boxtimes	$\bar{\boxtimes}$ or $\bar{\boxtimes}$
b	$\bar{\boxtimes}$	\boxtimes
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b	\boxtimes	$\bar{\boxtimes}$

$$P(\mathcal{Y} = \bar{\boxtimes} \text{ or } \bar{\boxtimes} | Y = \bar{\boxtimes}, X = g) = P(\mathcal{Y} = \bar{\boxtimes} \text{ or } \bar{\boxtimes} | Y = \bar{\boxtimes}, X = b)$$

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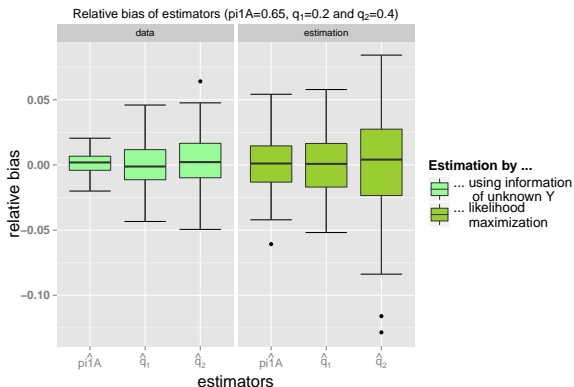
the coarsening mechanisms do not
depend on subgroup x

Case 2 - Binary covariates, no intercept

log-Likelihood involving binary covariate X :

$$\begin{aligned}
 l(\pi_{0A}, \pi_{1A}, q_1, q_2) &= n_{1A} \cdot [\ln(1 - q_1) + \ln(\pi_{1A})] + n_{0A} \cdot [\ln(1 - q_1) + \ln(\pi_{0A})] \\
 &+ n_{1B} \cdot [\ln(1 - q_2) + \ln(1 - \pi_{1A})] + n_{0B} \cdot [\ln(1 - q_2) + \ln(1 - \pi_{0A})] \\
 &+ n_{1AB} \cdot [q_1 \pi_{1A} + q_2(1 - \pi_{1A})] + n_{0AB} \cdot [q_1 \pi_{0A} + q_2(1 - \pi_{0A})]
 \end{aligned}$$

with $\pi_{01} = \frac{1}{2}$



Case 2 - Binary covariates, no intercept

Questions / Discussion suggestions:

- Is this result reasonable? (Remember: The coarsening mechanism does not depend on the values of X)

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- Is it possible to derive those estimators by means of equations like in the *iid* case, but now for each subgroup of X ?

$$\frac{n_{1AB}}{n_1} = \hat{q}_1 \cdot \hat{\pi}_{1A} + \hat{q}_2 \cdot (1 - \hat{\pi}_{1A})$$

$$\frac{n_{0AB}}{n_0} = \hat{q}_1 \cdot \hat{\pi}_{0A} + \hat{q}_2 \cdot (1 - \hat{\pi}_{0A})$$

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 \Rightarrow

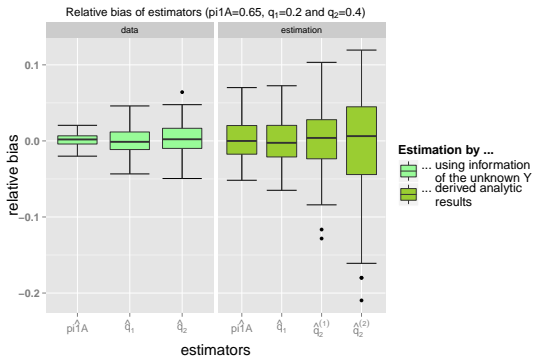
Resulting estimators

$$\hat{\pi}_{1A} = \frac{n_{1A} \cdot n_0}{2 \cdot n_1 \cdot n_{0A}}$$

$$\hat{q}_1 = 1 - 2 \cdot \frac{n_{0A}}{n_0}$$

and $\hat{q}_2^{(1)}$ and $\hat{q}_2^{(2)}$

Case 2 - Binary covariates, no intercept







$$\hat{q}_2^{(1)} = \frac{2 \cdot n_{0AB} - n_0 + 2 \cdot n_{0A}}{n_0}$$

$$\hat{q}_2^{(2)} = \frac{\frac{n_{1AB}}{n_1} - \frac{n_{1A}(n_0 - 2n_{0A})}{2n_1 n_{0A}}}{\frac{2n_1 n_{0A} - n_{1A} n_0}{2n_1 n_{0A}}}$$

Summary

- Important to distinguish between epistemic and ontologic uncertainty
- One can deal with ontologic uncertainty by redefining the sample space
- In case of ...
 - ... iid variables under epistemic uncertainty, a set of estimators results characterized by a special condition
 - ... being a binary covariate available, precise real valued point estimators seem to result

References

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